

On the existence of paths between points in high level excursion sets of Gaussian random fields

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Let $\mathbf{X} = (X(\mathbf{t}), \mathbf{t} \in \mathbb{R}^d)$ be a real-valued sample continuous Gaussian random field.

Given a level u , the excursion set of \mathbf{X} above the level u is the random set

$$A_u = \{\mathbf{t} \in \mathbb{R}^d : X(\mathbf{t}) > u\}.$$

Much is known about the structure of the excursion set when the field is smooth, and u is large (Adler and Taylor (2007), Azaïs and Wschebor (2009)).

Consider a large level u .

Question: given that two points in \mathbb{R}^d belong to the excursion set, **what is the probability that they belong to the same path connected component of the excursion set?**

Let $\mathbf{a}, \mathbf{b} \in \mathbb{R}^d$, $\mathbf{a} \neq \mathbf{b}$. A path in \mathbb{R}^d connecting \mathbf{a} and \mathbf{b} is a continuous map $\xi : [0, 1] \rightarrow \mathbb{R}^d$ with $\xi(0) = \mathbf{a}$, $\xi(1) = \mathbf{b}$.

Denote the collection of all such paths by $\mathcal{P}(\mathbf{a}, \mathbf{b})$. Estimate

$$P \left(\exists \xi \in \mathcal{P}(\mathbf{a}, \mathbf{b}) : X(\xi(v)) > u, 0 \leq v \leq 1 \mid X(\mathbf{a}) > u, X(\mathbf{b}) > u \right).$$

The non-trivial part of the problem: estimate the probability

$$\Psi_{\mathbf{a},\mathbf{b}}(u) := P(\exists \xi \in \mathcal{P}(\mathbf{a}, \mathbf{b}) : X(\xi(v)) > u, 0 \leq v \leq 1).$$

If the random field is stationary, we may assume that $\mathbf{b} = \mathbf{0}$, and use the notation $\Psi_{\mathbf{a}}$.

If the domain of a random field is restricted to $T \subset \mathbb{R}^d$, and \mathbf{a}, \mathbf{b} are in T , we consider

$$\Psi_{\mathbf{a},\mathbf{b}}(u) = P(\exists \xi \in \mathcal{P}(\mathbf{a}, \mathbf{b}) : \xi(v) \in T \text{ and } X(\xi(v)) > u, 0 \leq v \leq 1).$$

Large deviations setup

Let A be the open set

$$A \equiv A_{\mathbf{a}, \mathbf{b}} := \left\{ \omega \in C_0(\mathbb{R}^d) : \exists \xi \in \mathcal{P}(\mathbf{a}, \mathbf{b}), \omega(\xi(v)) > 1, 0 \leq v \leq 1 \right\},$$

$$C_0(\mathbb{R}^d) = \left\{ \omega = (\omega(\mathbf{t}), \mathbf{t} \in \mathbb{R}^d) \in C(\mathbb{R}^d) : \lim_{\|\mathbf{t}\| \rightarrow \infty} \omega(\mathbf{t}) / \|\mathbf{t}\| = 0 \right\}.$$

We can write for $u > 0$

$$\Psi_{\mathbf{a}, \mathbf{b}}(u) = P(u^{-1} \mathbf{X} \in A),$$

and use the large deviations results for Gaussian measures of Deuschel and Stroock (1989).

The reproducing kernel Hilbert space (RKHS) \mathcal{H} of the random field \mathbf{X} , is a subspace of $C(\mathbb{R}^d)$.

Consider the space of finite linear combinations $\sum_{j=1}^k a_j X(\mathbf{t}_j)$
 $a_j \in \mathbb{R}$, $\mathbf{t}_j \in \mathbb{R}^d$ for $j = 1, \dots, k$, $k = 1, 2, \dots$

Its closure \mathcal{L} in the mean square norm is identified with \mathcal{H} via the injection $\mathcal{L} \rightarrow C(\mathbb{R}^d)$ given by

$$H \rightarrow w_H = \left(E(X(\mathbf{t})H), \mathbf{t} \in \mathbb{R}^d \right)$$

and the resulting norm

$$\|w_H\|_{\mathcal{H}}^2 = E(H^2).$$

Let \mathbf{X} be stationary with spectral measure $F_{\mathbf{X}}$.

The RKHS \mathcal{H} can be identified with the subspace of $L^2(F_{\mathbf{X}})$ of functions with even real parts and odd imaginary parts, via the injection $L^2(F_{\mathbf{X}}) \rightarrow C_0(\mathbb{R}^d)$ given by

$$h \rightarrow S(h) = \left(\int_{\mathbb{R}^d} e^{i(\mathbf{t}, \mathbf{x})} \bar{h}(\mathbf{x}) F_{\mathbf{X}}(d\mathbf{x}), \mathbf{t} \in \mathbb{R}^d \right),$$

with the resulting norm

$$\|S(h)\|_{\mathcal{H}}^2 = \|h\|_{L^2(F_{\mathbf{X}})}^2 = \int_{\mathbb{R}^d} \|h(x)\|^2 F_{\mathbf{X}}(d\mathbf{x}).$$

Theorem 1 Let $\mathbf{X} = (X(\mathbf{t}), \mathbf{t} \in \mathbb{R}^d)$ be a continuous stationary Gaussian random field, with covariance function satisfying

$$\limsup_{\|\mathbf{t}\| \rightarrow \infty} R_{\mathbf{X}}(\mathbf{t}) \leq 0.$$

Then

$$\lim_{u \rightarrow \infty} \frac{1}{u^2} \log \Psi_{\mathbf{a}}(u) = -\frac{1}{2} \mathcal{C}_{\mathbf{X}}(\mathbf{a}),$$

where

$$\mathcal{C}_{\mathbf{X}}(\mathbf{a}) := \inf \left\{ \int_{\mathbb{R}^d} \|h(\mathbf{x})\|^2 F_{\mathbf{X}}(d\mathbf{x}) : \text{for some } \xi \in \mathcal{P}(\mathbf{0}, \mathbf{a}) \right. \\ \left. \int_{\mathbb{R}^d} e^{i(\xi(v), \mathbf{x})} \bar{h}(\mathbf{x}) F_{\mathbf{X}}(d\mathbf{x}) > 1, 0 \leq v \leq 1 \right\}.$$

The constraints in the optimization problem in Theorem 1 are not convex. However, **for a fixed path**, the constraints are convex, and one can use the convex Lagrange duality.

Theorem 2 For a continuous stationary Gaussian random field \mathbf{X} ,

$$\mathcal{C}_{\mathbf{X}}(\mathbf{a}) = \left[\sup_{\xi \in \mathcal{P}(\mathbf{0}, \mathbf{a})} \min_{\mu \in M_1^+([0,1])} \int_0^1 \int_0^1 R_{\mathbf{X}}(\xi(u) - \xi(v)) \mu(du) \mu(dv) \right]^{-1}.$$

Here $M_1^+([0,1])$ is the space of all probability measures on $[0,1]$.

An optimal path is a path of *maximal $R_{\mathbf{X}}$ capacity*.

Assume a path $\xi \in \mathcal{P}(\mathbf{0}, \mathbf{a})$ is fixed. Then the minimization problem

$$\min_{\mu \in M_1^+([0,1])} \int_0^1 \int_0^1 R_{\mathbf{X}}(\xi(u), \xi(v)) \mu(du) \mu(dv)$$

is the problem of finding a probability measures μ of *minimal energy*, or *capacitory measures*.

The set $\mathcal{W}_\xi \subseteq M_1^+([0,1])$ over which the minimum is achieved is a weakly compact convex subset of $M_1^+([0,1])$.

If the feasible set for the primary problem is non-empty then, for every $\varepsilon > 0$,

$$P \left(\sup_{0 \leq v \leq 1} \left| \frac{1}{u} X(\xi(v)) - x_\xi(v) \right| \geq \varepsilon \mid X(\xi(v)) > u, 0 \leq v \leq 1 \right) \rightarrow 0$$

as $u \rightarrow \infty$. Here H_ξ is primary optimal, and

$$x_\xi(v) = E[X(\xi(v))H_\xi], \quad 0 \leq v \leq 1.$$

Furthermore, there is a characterization of the optimal $H_\xi \in \mathcal{L}$ and the optimal $\mu \in \mathcal{W}_\xi$.

Theorem 3

(i) For every $\mu \in \mathcal{W}_\xi$ we have

$$H_\xi = \mathcal{C}_\mathbf{X}(\mathbf{a}, \mathbf{b}; \xi) \int_0^1 X(\xi(v)) \mu(dv)$$

with probability 1.

(ii) A probability measure $\mu \in M_1^+([0, 1])$ is a measure of minimal energy if and only if

$$\begin{aligned} & \min_{0 \leq v \leq 1} \int_0^1 R_\mathbf{X}(\xi(u), \xi(v)) \mu(du) \\ &= \int_0^1 \int_0^1 R_\mathbf{X}(\xi(u_1), \xi(u_2)) \mu(du_1) \mu(du_2) > 0. \end{aligned}$$

Remarks

- By Theorem 3, the function

$$v \mapsto \int_0^1 R_{\mathbf{X}}(\xi(u), \xi(v)) \mu(du), \quad 0 \leq v \leq 1,$$

is constant on the support of any measure $\mu \in \mathcal{W}_{\xi}$. This seems to indicate that the support of any measure of minimal energy may not be 'large'. However, this intuition holds only in some cases.

- If the random field is stationary, and the spectral measure is of the full support, then the image of any measure $\mu \in \mathcal{W}_{\xi}$ on the path ξ is unique.

One-dimensional case

In this case there is, essentially, a single path between two points.

Let $\mathbf{X} = (X(t), t \in \mathbb{R})$ be a stationary continuous Gaussian process. We are interested in understanding how the probability

$$\Psi_a(u) = P(X(t) > u, 0 \leq t \leq a)$$

changes with $a > 0$, and what happens with the optimal probability measures μ_a and limiting shapes x_a .

For some processes (in particular, those with a finite second spectral moment), on short intervals we get an easy description.

Proposition 1 Suppose that for some $a > 0$

$$R_{\mathbf{X}}(t) + R_{\mathbf{X}}(a - t) \geq R_{\mathbf{X}}(0) + R_{\mathbf{X}}(a) > 0 \text{ for all } 0 \leq t \leq a.$$

Then a measure in \mathcal{W}_a is given by

$$\mu = \frac{1}{2}\delta_0 + \frac{1}{2}\delta_1.$$

Furthermore,

$$C_{\mathbf{X}}(a) = \frac{2}{R_{\mathbf{X}}(0) + R_{\mathbf{X}}(a)},$$
$$x_a(t) = \frac{R_{\mathbf{X}}(t) + R_{\mathbf{X}}(a - t)}{R_{\mathbf{X}}(0) + R_{\mathbf{X}}(a)}, \quad 0 \leq t \leq a.$$

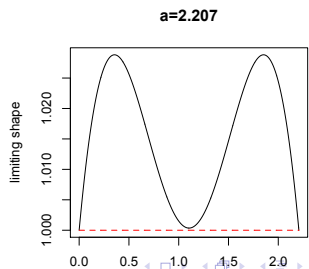
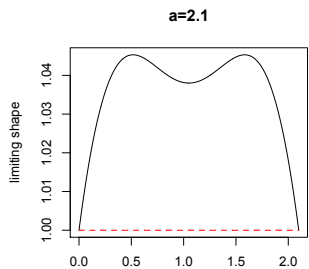
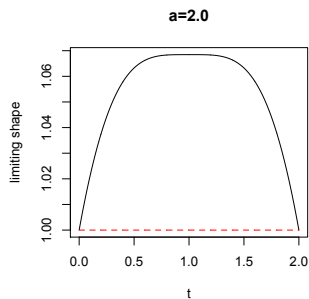
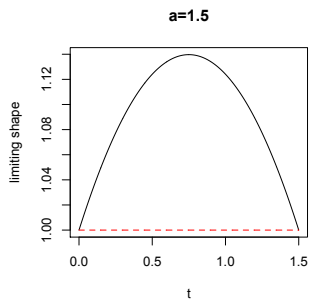
In any case, the measure $\mu = (\delta_0 + \delta_1)/2$ does NOT remain optimal for longer intervals.

Example 1 Consider the centered stationary Gaussian process with the Gaussian covariance function

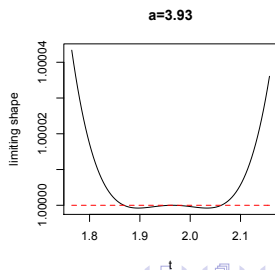
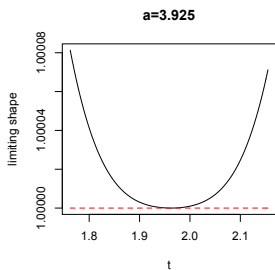
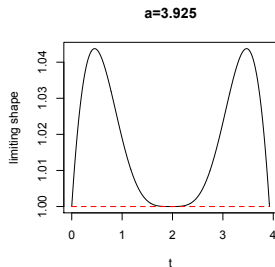
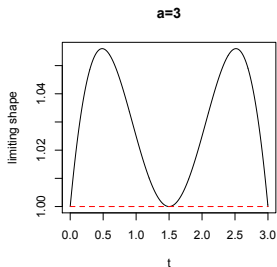
$$R(t) = e^{-t^2/2}, \quad t \in \mathbb{R}.$$

Since the spectral measure has a Gaussian spectral density which is of full support in \mathbb{R} , for every $a > 0$ there is a unique (symmetric) measure of minimal energy. Furthermore, the second spectral moment is finite, so that, for $a > 0$ sufficiently small this process satisfies the conditions of Proposition 1.

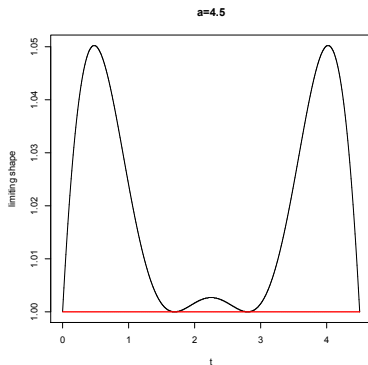
The measure $\mu = (\delta_0 + \delta_1)/2$ remains optimal for $a \leq a_1 \approx 2.2079$.



In the next regime the optimal measure acquires a point in the middle of the interval. This continues for $a_1 < a \leq a_2 \approx 3.9283$.



In the next regime the middle point of the optimal measure splits in two and starts moving away from the middle. This continues for $a_2 < a \leq a_3 \approx 5.4508$.



Example 2 Consider an Ornstein-Uhlenbeck process, i.e. a centered stationary Gaussian process with the covariance function

$$R(t) = e^{-|t|}, \quad t \in \mathbb{R}.$$

For this process the spectral measure has a Cauchy spectral density, which has a full support in \mathbb{R} . Therefore, for every $a > 0$ there is a unique (symmetric) measure of minimal energy.

In this case even the first spectral moment is infinite. Proposition 1 does not apply here.

The optimal probability measure is

$$\mu = \frac{1}{a+2}\delta_0 + \frac{1}{a+2}\delta_1 + \frac{a}{a+2}\lambda,$$

where λ is Lebesgue measure on $(0, 1)$.

There are no phase transitions. We have

$$C_{\mathbf{X}}(a) = (a+2)/2 \text{ for all } a > 0$$

and the limiting shape x_a is identically equal to 1 on $[0, a]$.

Long intervals

As the length a of the interval increases, a difference between short and long memory processes arises. In the short memory case, the uniform measure is asymptotically optimal.

Theorem 4 Assume that $R_{\mathbf{X}}$ is positive, and satisfies

$$\int_0^{\infty} R(t) dt < \infty.$$

Then, with λ denoting the uniform probability measure on $[0, 1]$,

$$\begin{aligned} \lim_{a \rightarrow \infty} \frac{1}{a} C_{\mathbf{X}}(a) &= \left(\lim_{a \rightarrow \infty} a \int_0^1 \int_0^1 R_{\mathbf{X}}(a(u-v)) \lambda(du) \lambda(dv) \right)^{-1} \\ &= \frac{1}{2 \int_0^{\infty} R(t) dt}. \end{aligned}$$

In the long memory case the uniform measure is no longer asymptotically optimal. We will assume that the covariance function of the process is regularly varying at infinity:

$$R_{\mathbf{X}}(t) = \frac{L(t)}{|t|^\beta}, \quad 0 < \beta < 1,$$

where L is slowly varying at infinity.

Consider the minimization problem with respect to Riesz kernel,

$$\min_{\mu \in M_1^+([0,1])} \int_0^1 \int_0^1 \frac{\mu(du)\mu(dv)}{|u-v|^\beta} \quad 0 < \beta < 1.$$

An optimal measure μ_β exists, but it is different from the uniform measure.

Theorem 5 Assume that $R_{\mathbf{X}}$ is positive and regularly varying. Then for any $\mu_{\beta} \in \mathcal{W}_{\beta}$, the set of optimal measures for the Riesz kernel,

$$\lim_{a \rightarrow \infty} R_{\mathbf{X}}(a) \mathcal{C}_{\mathbf{X}}(\mathbf{a}) = \left(\int_0^1 \int_0^1 \frac{\mu_{\beta}(du) \mu_{\beta}(dv)}{|u-v|^{\beta}} \right)^{-1}.$$

In particular, $\mathcal{C}_{\mathbf{X}}(\mathbf{a})$ is regularly varying with exponent β .