## **Credit Risk Modeling**

#### **References:**

- An Introduction to Credit Risk Modeling by Bluhm, Overbeck and Wagner, *Chapman & Hall*, 2003
- Credit Risk by Duffie and Singleton, New Age International Publishers, 2005
- Credit Risk Modeling and Valuation: An Introduction, by Kay Giesecke, http://www.stanford.edu/dept/MSandE/people/faculty/giesecke/introduction.pdf, 2004
- Options, Futures, and Other Derivatives, Hull, Prentice Hall India

#### The Basics of Credit Risk Management

- Loss Variable  $\tilde{L} = EAD \times SEV \times L$
- Exposure at Default (EAD) =  $OUTST + \gamma COMM$

**Basel Committee on banking supervision:** 75% of off-balance sheet amount. **Ex.** Committed line of one billion, current outstandings 600 million,  $EAD = 600 + 75\% \times 400 = 900.$ 

- Loss Given Default (LGD) = E[SEV]
  - Quality of collateral
  - Seniority of claim
- $L = 1_D$ , P(D) = DP: Probability of Default
  - Calibration from market data, Ex. KMV Corp.
  - Calibration from ratings, Ex. Moodys, S & P, Fitch, CRISIL : Statistical tools + Soft factors
  - Ratings  $\rightarrow$  DP: Fit "curve" to RR *vs* average DP plot

- Expected Loss (EL)  $E[\tilde{L}] = EAD \times LGD \times DP$
- Unexpected Loss (UL)  $= \sqrt{V(\tilde{L})}$

$$= EAD \times \sqrt{V(SEV) \times DP^2 + LGD^2 \times DP(1 - DP)}$$

**Portfolio:**  $\tilde{L}_{PF} = \sum_{i=1}^{m} EAD_i \times SEV_i \times L_i$ 

- $EL_{PF} = \sum_{i=1}^{m} EAD_i \times LGD_i \times DP_i$
- $UL_{PF}^2 = \sum_{i,j=1}^m EAD_i \times EAD_j \times Cov(SEV_i \times L_i, SEV_j \times L_j)$
- Constant Severities

$$=\sum_{i,j=1}^{m} EAD_i \times EAD_j \times LGD_i \times LGD_j \times \sqrt{DP_i(1-DP_i)DP_j(1-DP_j)} \rho_{ij}$$

• Value at Risk (VaR):  $q_{\alpha}$ 

$$q_{\alpha} : \inf\{q > 0 : P[L_{PF} \le q] \ge \alpha\}$$

- Economic Capital  $(EC_{\alpha}) = q_{\alpha} EL_{PF}$
- Expected Shortfall:

$$TCE_{\alpha} = E[\tilde{L}_{PF} \mid \tilde{L}_{PF \ge q_{\alpha}}]$$

• Economic Capital based on Shortfall Risk:  $TCE_{\alpha} - EL_{PF}$ 

- Loss Distribution
  - Monte-Carlo Simulation
  - Analytical Approximation: Credit Risk<sup>+</sup>
- Today's Industry Models
  - Credit Metrics and KMV-Model
  - Credit Risk<sup>+</sup>
  - CreditPortfolio View
  - Dynamic Intensity Models

## Credit Metrics and the KMV-Model

- Asset Price Process:  $A_t^{(i)}$
- Valuation Horizon: T

$$L_i = 1_{\{A_T^{(i)} < C_i\}} \sim B(1; P(A_T^{(i)} < C_i))$$

$$r_i = \log\left(\frac{A_T^{(i)}}{A_0^{(i)}}\right) = R_i \Phi_i + \epsilon_i, \qquad i = 1, 2, \dots, m$$

- Firm's composite factor Φ<sub>i</sub> is a superposition of systematic influences (industry and country indices)
- $\epsilon_i$  : Firm specific or idiosyncratic part
- $R_i^2 =$  portion of the volatility in  $r_i$  explained by volatility in  $\Phi_i$

$$r_i \sim N(0,1);$$
  $\Phi_i \sim N(0,1);$   $\epsilon_i \sim N(0,1-R_i^2)$ 

#### **Global Correlation Model**

**Industry and Country Indices:** 

$$\Psi_j = \sum_{n=1}^N b_{j,n} \Gamma_n + \delta_n, \qquad j = 1, \dots, J$$

Independent Global Facors:

$$\Phi_i = \sum_{j=1}^J w_{ij} \Psi_j$$

$$L_i = 1_{\{r_i < c_i\}} \sim B(1, P(r_i < c_i))$$

$$r_i < c_i \equiv \epsilon_i < c_i - R_i \Phi_i$$

 $\Gamma_n$ 

$$p_{i} = P(r_{i} < c_{i}) \implies c_{i} = N^{-1}(p_{i})$$

$$p_{i}(\Phi_{i}) = N \left[ \frac{N^{-1}(p_{i}) - R_{i}\Phi_{i}}{\sqrt{1 - R_{i}^{2}}} \right]$$

$$p_{i}(\Phi_{i}) = N \left[ \frac{N^{-1}(p_{i}) - R_{i}\sum_{j=1}^{J} w_{ij}\Psi_{j}}{\sqrt{1 - R_{i}^{2}}} \right]$$

- Simulate a realization of  $\Psi_j \longrightarrow$  Simulate realization of  $L_i$ 
  - $\rightarrow$  **One** realization of the loss
- Loss Distribution

- KMV tool GCorr computes asset correlations
- KMV provides the weights and asset correlations to its customers
- Can use these correlations with heavy tailed copulas to obtain stronger tail dependencies:
  - $F_n$  Univariate t-distribution with n d.f.
  - $F_{n,\Gamma}$  Multivariate t-distribution with n d.f. and correlation matrix  $\Gamma$ .

- 
$$C_{n,\Gamma}(u_1,\ldots,u_m) = F_{n,\Gamma}(F_n^{-1}(u_1),\ldots,F_n^{-1}(u_m))$$

$$- \Phi_{n,\Gamma}(x_1,\ldots,x_m) = C_{n,\Gamma}(N(x_1),\ldots,N(x_m))$$

#### **Two Differences Between KMV-Model and Credit Metrics**

- Credit Metrics uses equity price correlations, whereas KMV carries out the complicated translation from equity and market information to asset values
- Credit Metrics uses indices referring to a combination of some industry in some particular country, whereas KMV considers industries and countries separately

# CreditPortfolio View

- **Default and rating migrations** are subject to random fluctuations that depend on the **economic cycle**
- Unconditional migration matrix  $\bar{M} = (\bar{m}_{ij}), \quad i, j = 1, \dots, K$ : rating categories
- $\bar{m}_{iK}$ : one year historic probability of default in rating category i
- $\bullet~S$  risk segments that react differently to the economic conditions

- 1. Simulate a segment specific conditional default probability  $p_s$ , s = 1, ..., S. Aggregated Second Level Scenario
- 2. Define the risk index

$$r_s = \frac{p_s}{\bar{p}_s}$$

- $\bar{p}_s$  unconditional default probability of segment s
- 3. Conditional migration matrix  $M^{(s)}$ :

$$m_{ij}^s = \alpha_{ij}(r_s - 1) + \bar{m}_{ij}$$

The shift matrix  $(\alpha_{ij})$  satisfying  $\sum_j \alpha_{ij} = 0$  must be calibrated by the user

$$\alpha_{ij} \ge 0, \qquad i < j, \qquad \alpha_{ij} \le 0, \qquad i > j$$

 $M^{(s)}$  applies to all obligors in segment s. Some entries may turn out to be negative. Set equal to 0 and renormalize.

$$m_{ij}^s = \alpha_{ij}(r_s - 1) + \bar{m}_{ij}$$

- $r_s < 1$ : expanding economy, lower possibility of downgrades and higher number of upgrades
- $r_s = 1$ : average macroeconomic scenario
- $r_s > 1$ : recession, downgrades more likely

For each realization of the default probabilities, simulate the defaults and loss. Repeat simulation several times to generate the loss distribution. **CPV** supports two modes of calibration:

• CPV Macro: default and rating migrations are explained by a macroeconomic regression model. Macroeconomic model is calibrated by means of times series of empirical data.

$$Y_{s,t} = w_{s,0} + \sum_{j=1}^{M} w_{s,j} X_{s,j,t} + \epsilon_{s,t}, \qquad \epsilon_{s,t} \sim N(0, \sigma_{s,t}^2)$$

$$X_{s,j,t} = \theta_{j,0} + \sum_{k=1}^{t_0} \theta_{j,k} X_{s,j,t-k} + \eta_{s,j,t}$$
$$p_{s,t} = \frac{1}{1 + \exp(-y_{s,t})}$$

• CPV Direct:  $p_s$  drawn from a gamma distribution. Need only to calibrate the two parameters of the gamma distribution for each s.  $p_s$  can turn out to be larger than 1.

### **Dynamic Intensity Models**

• Basic Affine or Intensity Process

$$d\lambda(t) = \kappa(\theta - \lambda(t)) dt + \sigma \sqrt{\lambda(t)} dB(t) + \Delta J(t)$$

- J(t): pure jump process independent of the BM B(t) with jumps arriving according to a Poisson process with rate  $\ell$  and jump sizes  $\Delta J(t) \sim \exp(\mu)$
- $\kappa =$  mean-reversion rate;  $\sigma =$  diffusive volatility;  $\bar{m} = \theta + \ell \mu / \kappa$  long-run mean
- Unconditional Default Probability  $q(t) = E\left[e^{-\int_0^t \lambda(u) \ du}\right]$
- Correlated defaults  $\lambda_i = X_c + X_i$

 $X_c, X_i$  basic affine processes with parameters  $(\kappa, \theta_c, \sigma, \mu, \ell_c)$  and  $(\kappa, \theta_i, \sigma, \mu, \ell_i)$ representing the **common performance aspects** and the **idiosyncratic risk** 

•  $\lambda_i$ : basic affine process with parameters  $(\kappa, \theta_c + \theta_i, \sigma, \mu, \ell_c + \ell_i)$ 

$$dX_p(t) = \kappa(\theta_p - X_p(t)) dt + \sigma \sqrt{X_p(t)} dB^p(t) + \Delta J^p(t), \qquad p = c, a$$

$$d(X_c + X_i)(t) = \kappa((\theta_c + \theta_i) - (X_c + X_i)(t)) dt + \sigma(\sqrt{X_c(t)} dB^c(t) + \sqrt{X_i(t)} dB^i(t)) + \Delta(J^c + J^i)(t)$$

$$dW_t = \sqrt{\frac{X_t^c}{X_t^c + X_t^i}} \, dB^c(t) + \sqrt{\frac{X_t^i}{X_t^c + X_t^i}} \, dB^i(t)$$
$$d(X_c + X_i)(t) = \kappa((\theta_c + \theta_i) - (X_c + X_i)(t)) \, dt + \sigma \sqrt{(X^c + X^i)(t)} \, dW(t)$$
$$+ \Delta (J^c + J^i)(t)$$

Conditioned on a realization of  $\lambda_i(t)$ ,  $0 \le t \le T$ , the default time of obligor *i* is the first arrival in a non-homogenous Poisson process with rate  $\lambda_i(\cdot)$ Conditional Probability of No Default  $= \exp(-\int_0^T \lambda(s) ds)$ 

## The Credit Risk<sup>+</sup> Model

- Introduced in 1997 by CSFB
- Actuarial Model
- One of the most widely used credit portfolio models

#### • Advantages:

- Loss Distribution can be computed analytically
- Requires no Monte-Carlo Simulations
- Explicit Formulas for Obligor Risk Contributions
- Numerically stable computational procedure (Giese, 2003)

### The Standard CR<sup>+</sup> Model

- Choose a suitable basic unit of currency  $\Delta L$
- Adjusted exposure of obligor A,  $\nu_A = \lfloor E_A / \Delta L \rfloor$
- Smaller number of **Exposure Bands**
- $p_A$  expected default probability
- The total portfolio loss  $L = \sum_{A} \nu_A N_A.$
- $N_A \in Z_+$  Default of obligor A
- PGF of the Loss Distribution  $G(z) = \sum_{n=0}^{\infty} P(L = n) z^{n}$ .

- Apportion Obligor Risk among K Sectors (Industry, Country) by choosing numbers  $g_k^A$  such that  $\sum_{k=1}^K g_k^A = 1$ .
- Sectoral Default Rates represented by non-negative variables  $\gamma_k$

$$E(\gamma_k) = 1,$$
  $Cov(\gamma_k, \gamma_l) = \sigma_{kl}$   $k = 1, ..., K.$ 

- Standard CR<sup>+</sup> Model assumes  $\sigma_{kl} = 0, \qquad k \neq l$
- Relating Obligor default rates to sectoral default rates

$$p_A(\gamma) = p_A \sum_{k=1}^K g_k^A \gamma_k,$$

- $p_A(\gamma)$  default rate conditional on the sector default rates  $\gamma = (\gamma_1, \ldots, \gamma_K)$ .
- Specific Sector:  $\gamma_0 \equiv 1$ . Captures Idosyncratic Risk

- Conditional on  $\gamma$  default variables  $N_A$  assumed to be independent Poisson
- Main Criticism of CR<sup>+</sup> Model. Not Fair

$$-p_A = 0.1 \longrightarrow P(N_A = 2) = 0.0045$$

- Need not assume  $N_A$  is Poisson, but Bernoulli
- Conditional PGF

$$G_{\gamma}(z) = \exp(\sum_{k=1}^{K} \gamma_k P_k(z)),$$

$$P_{k}(z) = \sum_{A} g_{k}^{A} p_{A}(z^{\nu_{A}} - 1)$$
$$= \sum_{m=1}^{M} \left( \sum_{\{\nu_{A}=m\}} g_{k}^{A} p_{A} \right) (z^{m} - 1)$$

• Number of defaults in any exposure band is Poisson

- **Default correlation** between obligors arise only through their dependence on the common set of sector default rates
- Unconditional PGF of Loss Distribution

$$G(z) = E^{\gamma}(\exp(\sum_{k=1}^{K} \gamma_k P_k(z))) = M_{\gamma}(T = P(z))$$

• MGF of Univariate Gamma Distribution with Mean 1 and Variance  $\sigma_{kk}$  is  $(1 - \sigma_{kk}t_k)^{-\frac{1}{\sigma_{kk}}}$ 

$$G^{CR+}(z) = \exp\left(-\sum_{k=1}^{K} \frac{1}{\sigma_{kk}} \log(1 - \sigma_{kk} P_k(z))\right)$$

• Giese(2003): Numerically Stable Fast Recursion Scheme

## The Compound Gamma CR<sup>+</sup> Model (Giese, 2003)

- $\bullet\,$  Introduce sectoral correlations via common scaling factor S
- Conditional on  $S \gamma_K$  is Gamma distributed with shape parameter  $\hat{\alpha}_k(S) = S \alpha_k, \qquad \alpha_k > 0$ , and constant scale parameter  $\beta_k$ .
- S follows Gamma with E[S] = 1 and  $Var(S) = \hat{\sigma}^2$ .
- $1 = E\gamma_k = \alpha_k\beta_k$
- $\sigma_{kl} = \delta_{kl}\beta_k + \hat{\sigma}^2$
- Uniform Level of Cross Covariance ⇒ Distortion of Correlation Structure.

$$M_{\gamma}^{CG}(T) = \exp\left\{-\frac{1}{\hat{\sigma}^2}\log\left[1 + \hat{\sigma}^2\sum_{k=1}^{K}\frac{1}{\beta_k}\log(1 - \beta_k t_k)\right]\right\}$$

• Calibration Problems

# The Two Stage CR<sup>+</sup> Model (SKI, AD)

•  $Y_1, \ldots, Y_N$ : Common set of Uncorrelated Risk Drivers

$$\gamma_k = \sum_{i=1}^N a_{ki} Y_i$$

- $Y_i \sim \text{Gamma}$  with mean 1 and variance  $V_{ii}$
- Principle Component Analysis of Macroeconomic Variables
- Factor Analysis

$$G(z) = E^{\gamma}(exp(\sum_{k=1}^{K} \gamma_k P_k(z))) = E^{Y}(exp(\sum_{k=1}^{K} (\sum_{i=1}^{N} a_{ki} Y_i) P_k(z)))$$
  
=  $E^{Y}(exp(\sum_{i=1}^{N} (\sum_{k=1}^{K} a_{ki} P_k(z)) Y_i))$   
=  $E^{Y}(exp(\sum_{i=1}^{N} Y_i Q_i(z))) = M_Y(T = Q(z))$ 

$$Q_i(z) = \sum_{k=1}^K a_{ki} P_k(z)$$

$$G(z) = \exp\left\{-\sum_{i=1}^{N} \frac{1}{\sigma_{ii}} \log(1 - \sigma_{ii}Q_i(z))\right\}$$

#### **Model Comparison**

- Giese (2003) had pointed out deficiencies in the earlier attempt to incorporate correlations due to Burgisser *et al*
- We compare the compound gamma and the two stage gamma models
- Test portfolio made up of K = 12 sectors, each containing 3,000 obligors
- Obligors in sectors 1 to 10 belong in equal parts to one of three classes with adjusted exposures  $E_1 = 1$ ,  $E_2 = 2.5$ , and  $E_3 = 5$  monetary units and respective default probabilities  $p_1 = 5.5\%$ ,  $p_2 = .8\%$ ,  $p_3 = .2\%$ .
- For the three obligor classes in sectors 11 and 12, we assume the same default rates but twice as large exposures ( $E_1 = 2$ ,  $E_2 = 5$ ,  $E_3 = 10$ )
- $\sigma_{kk} = 0.04, k = 1, \dots, 10$   $\sigma_{11,11} = \sigma_{12,12} = 0.49$
- Correlation between sectors 11 and 12 is 0.5 whereas correlations between all the other sectors are set equal to 0

• 
$$\gamma_i = Y_i, i = 1, ..., 11, \gamma_{12} = 0.5(Y_{11} + Y_{12}), \text{ with } Var(Y_{11}) = 0.49$$
  
 $Var(Y_{12}) = 1.47, \text{ and } Var(Y_i) = 0.04 \text{ for } i = 1, ..., 10$ 

	Standard CR+	Compound Gamma Model	Two-Stage Model
Expected Loss	1%	1%	1%
Std Deviation	0.15%	0.17%	0.17%
99% Quantile	1.42%	1.48%	1.53%
99.5% Quantile	1.49%	1.55%	1.62%
99.9% Quantile	1.64%	1.71%	1.84%

**Table 1:** Comparison of the loss distributions from the standard CR<sup>+</sup>, compound gamma and two stage models for the test portfolio in example 1. All loss statistics are quoted as percentage of the total adjusted exposure.

•  $\hat{\sigma}^2 = 0.013$ . This translates to a much lower correlation of 0.0265 (instead of 0.5) between sectors 11 and 12

## **Risk Contributions**

- Value at Risk VAR  $\ell_q$
- Economic Capital  $\ell_q E[L]$
- Expected Shortfall  $E[L|L \ge \ell_q]$
- Quantile Contribution  $QC_A$

$$QC_A = \nu_A E(N_A | L = \ell_q) = p_A \nu_A \frac{\sum_{k=1}^K g_A^k D^{(\ell_q - \nu_A)} G_k(z)}{D^{(\ell_q)} G(z)}$$
$$G_k(z) = \frac{\partial}{\partial t_k} M_Y(T = Q(z))$$

$$G_{k}(z) = G(z) \sum_{i=1}^{N} \frac{a_{ki}}{1 - \sigma_{ii}Q_{i}(z)}$$
  
=  $G(z) \left(\sum_{i=1}^{N} a_{k,i} exp(-log(1 - \sigma_{ii}Q_{i}(z)))\right).$ 

Sector	$CR^+$	Compound Gamma Model	Two-Stage Model	
1,2	24.25%	21.71%	27.42%	
$3,\ldots,10$	0.37%	1.64 %	0.2 %	
11, 12	24.25 %	21.71%	21.59 %	

**Table 2:** Aggregated risk contributions (in percent). Contributions to the loss variance for the risk-adjusted breakdown of VaR (on a 99.9% confidence level).

• Compound gamma model can't pick up differing correlations among sectors that are otherwise similar.